Ancestral processes with selection

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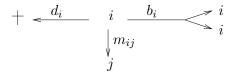
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mutation-selection differential equation & multitype branching process

- forward and backward
- ancestral distribution, variational principle (with H.-O. Georgii, 2007)
- 2 Moran model with selection and mutation
 - forward and backward
 - ancestral distribution, lookdown ancestral selection graph (with A. Wakolbinger, S. Kluth, U. Lenz, 2015, 2016, 2018)

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individual = (geno)type $i \in S$ (finite)



(Malthusian) fitness: $r_i := b_i - d_i$

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 $y_i(t)$ abundance of type i at time t $(i \in S)$

$$\dot{y}_i(t) = r_i y_i(t) + \sum_{j:j \neq i} (y_j(t)m_{ji} - y_i(t)m_{ij})$$

or

$$\dot{y}(t) = y(t) \left(\overbrace{\mathcal{M} + \mathcal{R}}^{\mathcal{A}} \right) \\ \uparrow_{m_{ij}} \uparrow_{r_i}^{\uparrow}$$

with solution

 $y(t) = y(0) e^{t\mathcal{A}}$

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relative frequencies:

$$\dot{p}_i(t) = \left(r_i - \overline{\boldsymbol{r}}(t)\right) p_i(t) + \sum_{j:j \neq i} \left(p_j(t)m_{ji} - p_i(t)m_{ij}\right)$$

$$p_i(t) := \frac{y_i(t)}{\sum_j y_j(t)}, \qquad \bar{r}(t) := \sum_j r_j p_j(t)$$

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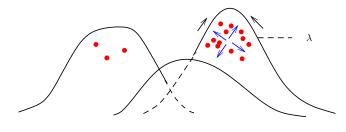
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 \mathcal{M} irreducible \rightsquigarrow Perron-Frobenius: $p\mathcal{A} = \lambda p \quad (\langle p, 1 \rangle = 1)$

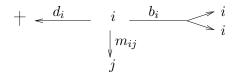
 $p(t) \stackrel{t \to \infty}{\longrightarrow} p$ (stationary type distribution)

 $\lambda = \sum_i r_i p_i = \langle p, r \rangle = ar{r}$ (equilibrium mean fitness)



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Multitype Branching



i-individual:

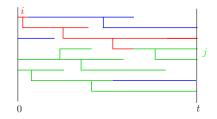
waiting time $\sim \mathcal{E}(a_i)$, $a_i = b_i + d_i + \sum_{j: j \neq i} m_{ij}$

then: birth, death, mutation to j with probability $\frac{b_i}{a_i}, \frac{d_i}{a_i}, \frac{m_{ij}}{a_i}$

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Multitype branching



(Z(t)), Z(t) counting measure on S $Z_j(t) \#$ ind. of type j at time tfirst-moment generator: $\mathcal{A} = \mathcal{M} + \mathcal{R}$

$$\begin{split} \mathbb{E}^i(Z_j(t)) &= (\,\mathrm{e}^{t\mathcal{A}})_{ij} \\ \text{assumption:} \quad \lambda > 0 \, \rightsquigarrow \, \text{branching supercritical} \\ & (\text{with asymptotic growth rate } \lambda \,) \end{split}$$

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Connections "branching" \leftrightarrow "MuSe":

$$\begin{array}{c|c} \mathbf{1} & \frac{Z(t)}{|Z(t)|} \stackrel{t \to \infty}{\stackrel{\frown}{\uparrow}} p \stackrel{\infty \leftarrow t}{\leftarrow} p(t) \\ & \text{a.s., on \{non-extinction\} (Kesten-Stigum '66)} \\ p \text{ left PF-EV of } \mathcal{A} \text{ (stationary distribution of types)} \end{array}$$

$$\mathbb{2} \mathbb{E}^{i}(|Z(t)| e^{-\lambda t}) \xrightarrow{t \to \infty} h_{i} \xleftarrow{\infty \leftarrow t} \frac{\sum_{j} (e^{t\mathcal{A}})_{ij}}{|p(0) e^{t\mathcal{A}}|}$$

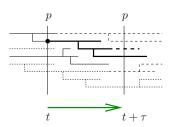
 $h=(h_i)_{i\in S}$ right PF-EV of ${\mathcal A}$ (asymptotic offspring expectation)

$$\langle p,1\rangle = 1 = \langle p,h\rangle$$

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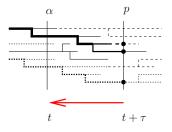
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Forward and backward



forward

backward (NO coalescent)



 $au o \infty \rightsquigarrow \ lpha_i := p_i h_i$ ancestral type distribution Jagers, Nerman 1992 . . .

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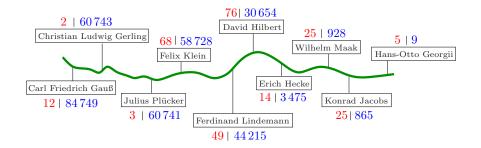
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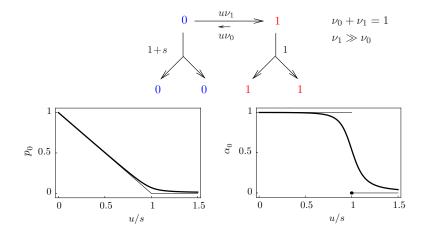
Mathematics genealogy project



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Forward and backward



 $s=0.001\text{,}~\nu_0=0.005$ and $\nu_0\rightarrow 0$

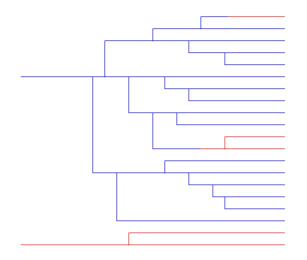
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Forward and backward



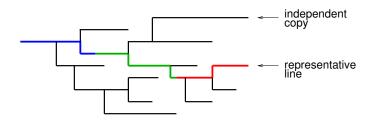
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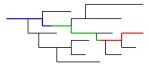
Large deviations



mutation process (M_t) on representative line $(M_t$ type at time t) (generator \mathcal{M} , stationary distribution π)

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empirical measure L_t on S: $L_t(j) = \frac{1}{t} \int_0^t \mathbf{1}_{\{M_\tau = j\}} d\tau$ (random!)

LDP:
$$\mathbb{P}(L_t \sim \nu) \approx e^{-t I_{\mathcal{M}}(\nu)} \quad (\text{large } t) \\ \left(\lim_{t \to \infty} \frac{1}{t} \log \mathbb{P}(L_t \in A) = -\inf_{\nu \in A} I_{\mathcal{M}}(\nu)\right)$$

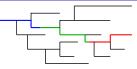
rate function: $I_{\mathcal{M}}(\nu) = \sup_{v>0} \left(-\langle \nu, \frac{\mathcal{M}v}{v} \rangle \right)$ (M_t) reversible $\rightsquigarrow I_{\mathcal{M}}(\nu) = -\left\langle \sqrt{\frac{\nu}{\pi}}, \mathcal{M}\sqrt{\frac{\nu}{\pi}} \right\rangle_{\pi}$

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Variational principle

line with $L_t = \nu$

experiences



• mutation: changes ν $\mathbb{P}(L_t \sim \nu)$ decays with $I_{\mathcal{M}}(\nu)$ (> 0 for $\nu \neq \pi$ (stat. distr.), = 0 for $\nu = \pi$) $L_t \xrightarrow[t \to \infty]{} \pi$ a.s. reproduction: duplicates ν at rate r_{M_t} at time tmean rate $\langle \nu, r \rangle$

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Theorem (EB & Georgii 2007)

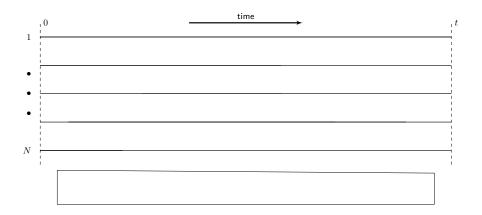
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2-type Moran model with mutation and selection

- \blacksquare population of fixed size N
- types: 0 ('fit') and 1 ('unfit')
- individuals of type 1 reproduce at rate 1
- individuals of type 0 reproduce at rate $1 + s^N$, $s^N \ge 0$
- single offspring inherits parent's type and replaces uniformly chosen individual
- mutation at rate $u^N > 0$
- resulting type: 0 with probability $\nu_0;\, 1$ with probability ν_1 $(\nu_0+\nu_1=1)$

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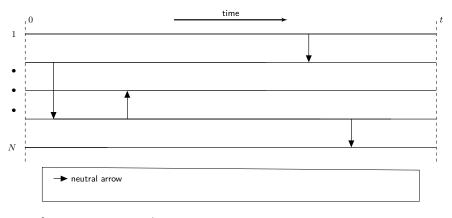


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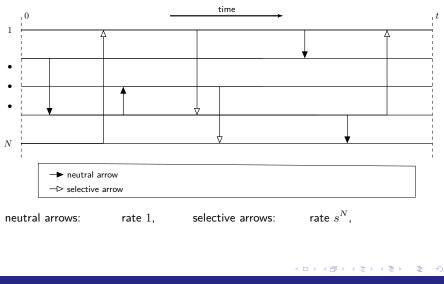


neutral arrows: rate 1,

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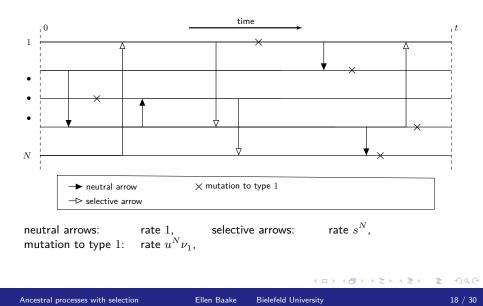
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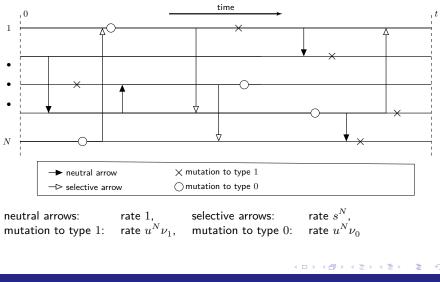
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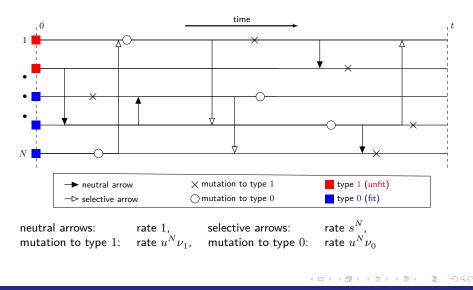
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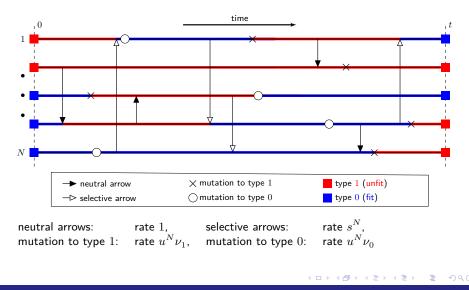
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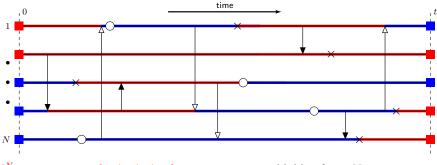
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 $Y_t^N :=$ proportion of individuals of type 1 at time t in MoMo of size N

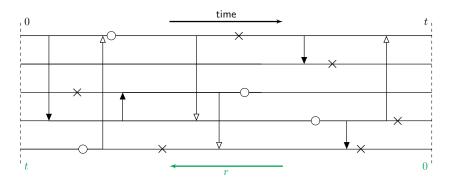
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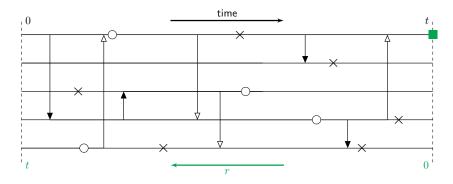


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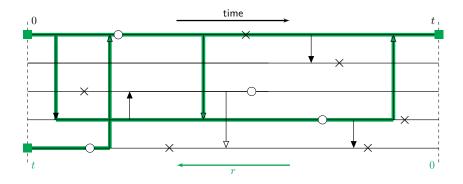


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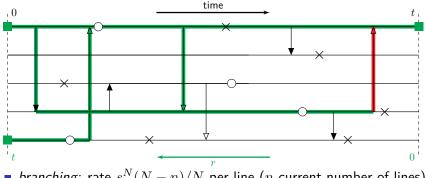


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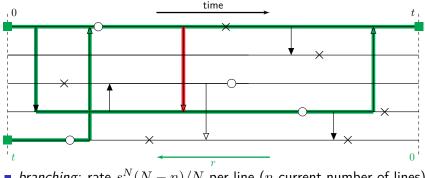
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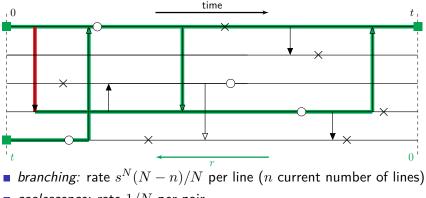
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branching: rate $s^N(N-n)/N$ per line (*n* current number of lines)

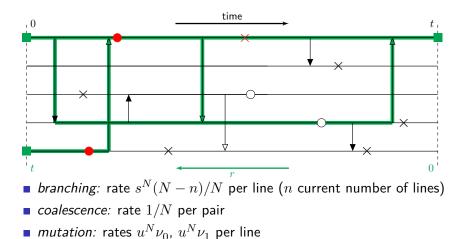


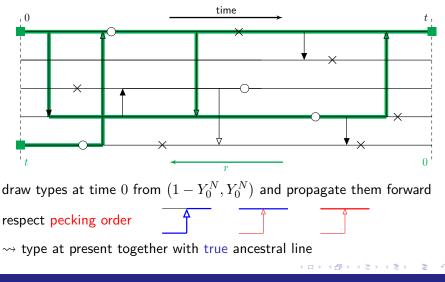
branching: rate $s^N(N-n)/N$ per line (*n* current number of lines)



coalescence: rate 1/N per pair

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 $N \to \infty$ s.t. $Ns^N \to \sigma$, $Nu^N \to \theta$, $Y_0^N \to y_0$ $\rightsquigarrow (Y_{tN}^N) \xrightarrow{d}$ Wright-Fisher diffusion (Y_t)

$$\begin{split} \mathrm{d}Y_t &= \sqrt{Y_t(1-Y_t) \,\mathrm{d}W_t} - \sigma Y_t(1-Y_t) \,\mathrm{d}t + (1-Y_t)\theta\nu_1 \,\mathrm{d}t - Y_t\theta\nu_0 \,\mathrm{d}t, \\ Y_0 &= y_0, \quad (W_t) \text{ standard Brownian motion} \end{split}$$

$$\theta > 0, \ 0 < \nu_0 < 1, t \to \infty \rightsquigarrow Y_t \xrightarrow{d} \widetilde{Y}$$

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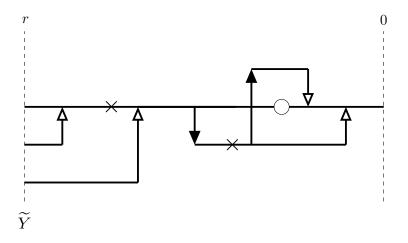
• branching at rate σ per line

• coalescence events at rate 1 per pair

 \blacksquare mutation superimposed on lines at rate $\theta \nu_0$ and $\theta \nu_1$

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ASG in diffusion limit



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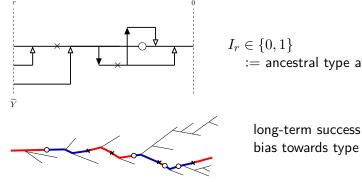
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Ancestral type



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 $r \to \infty \rightsquigarrow I_r \xrightarrow{d} \tilde{I}$

bias towards type 0

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$$\mathbb{P}(\tilde{I}=0), \, \mathbb{P}(\tilde{I}=1)?$$

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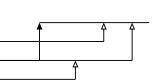
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arrange lines according to pecking order (exchangeability!)

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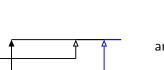
arrange lines according to pecking order (exchangeability!)



ancestral line is

- Iowest type-0 line if there is one
- immune line otherwise

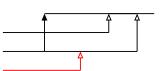
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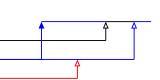
immune line otherwise

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arrange lines according to pecking order (exchangeability!)



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immune line otherwise

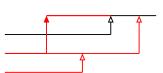
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arrange lines according to pecking order (exchangeability!)



ancestral line is

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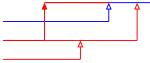
immune line otherwise

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arrange lines according to pecking order (exchangeability!)





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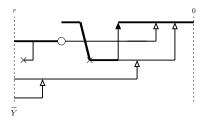
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Pruning upon mutation





ancestral line is

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■ immune line otherwise

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The line-counting process of the pruned ASG

- $L_r =$ number of lines at time r
- (L_r) Markov chain in continuous time with rates from



stationary distribution $(r \to \infty)$:

 $a_n := \mathbb{P}(\tilde{L} > n), \quad n \ge 0$

first-step analysis \rightsquigarrow (Fearnhead's) recursion:

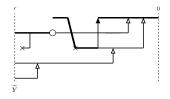
$$(n+1+\theta+\sigma)a_n=(n+1+\theta\nu_1)a_{n+1}+\sigma a_{n-1},\quad n>0,$$

$$a_0 = 1, \quad \lim_{n \to \infty} a_n = 0.$$

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Type of ancestral line



Theorem (Lenz, Kluth, EB, Wakolbinger 2015)

The type distribution of the ancestral line in the distant past is given by

$$\mathbb{P}(\tilde{I}=1\mid \tilde{Y}) = \sum_{n>0} (a_{n-1}-a_n) \tilde{Y}^n$$

$$\mathbb{P}(\tilde{I}=1) = \sum_{n>0} (a_{n-1} - a_n) b_n, \quad \mathbb{P}(\tilde{I}=0) = \sum_{n \ge 0} a_n (b_n - b_{n+1}),$$

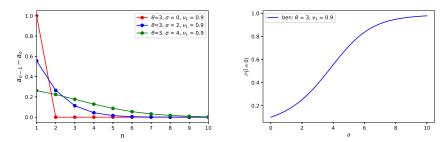
where $b_n := \mathbb{E}(\widetilde{Y}^n)$.

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distribution of $\widetilde{\boldsymbol{L}}$

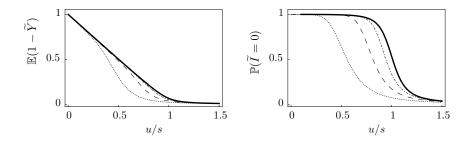
probability of fit ancestor

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Forward and backward



$$s=0.001,\ \nu_0=0.005$$

$$\sigma=Ns,\ \vartheta=Nu \ \text{with}\ N=10^4,\ N=3\cdot 10^4,\ N=10^5$$

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Large devations and variational principle for (multitype) MoMo?

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2